# Identifying Collocations to Measure Compositionality : Shared Task System Description

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## Abstract

This paper describes three systems from the University of Minnesota, Duluth that participated in the DiSCo 2011 shared task that evaluated distributional methods of measuring semantic compositionality. All three systems approached this as a problem of collocation identification, where strong collocates are assumed to be minimally compositional. duluth-1 relies on the t-score, whereas duluth-2 and duluth-3 rely on Pointwise Mutual Information (pmi). duluth-1 was the *top ranked system overall* in coarse–grained scoring, which was a 3-way category assignment where pairs were assigned values of high, medium, or low compositionality.

## 1 Introduction

An ngram or phrase that means more than the sum of its parts is said to be non-compositional. Well known examples include *kick the bucket* (i.e., to die) and *red tape* (i.e., bureaucratic steps). The ability to measure the degree of semantic compositionality in a unit of text is a key capability of NLP systems, since non-compositional phrases can be treated as a single unit, rather than as a series of individual words. This has a tremendous impact on word sense disambiguation systems, for example, since a noncompositional phrase will often have just one possible sense and thereby be reduced to a trivial case, whereas the combination of possible sense assignments for the words that make up a phrase can grow exponentially.

Identifying collocations is another key capability of NLP systems. Collocations are generally considered to be units of text that occur with some regularity and may have some non-compositional meaning. The Duluth systems that participated in the DiSCo 2011 shared task (Biemann and Giesbrecht, 2011) seek to determine the degree to which collocation identification techniques can be used to measure semantic compositionality. In particular, these systems are based on the following hypothesis:

An ngram that has a high score according to a measure of association (for identifying collocations) will be less compositional (and less literal) than those that have lower scores.

The intuition underlying this hypothesis is a high score from a measure of association shows that the words in the ngram are occurring together more often than would be expected by chance, and that a non-compositional phrase is unlikely to occur in such a way that it looks like a chance event.

## 2 System Development

The Duluth systems were developed by identifying collocations based on frequency counts obtained from the WaCky English corpus (Baroni et al., 2009), hereafter referred to as *the corpus*. The part of speech tags were removed from the corpus, and the text was converted to lower case. A set of 139 training pairs was provided by the task organizers that had been manually rated for compositionality. This gold standard data was used to select which measures of association would form the basis of the Duluth systems. Thereafter a separate set of 174 test pairs were provided by the organizers for evaluation.

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#### 2.1 Collocation Discovery

The Ngram Statistics Package (Text::NSP) (Banerjee and Pedersen, 2003) was used to measure the association between the training pairs based on frequency count data collected from the corpus. All thirteen measures in the Ngram Statistics Package were employed, including the Log-likelihood Ratio (II) (Dunning, 1993), Pointwise Mutual Information (pmi) (Church and Hanks, 1990), Mutual Information (tmi) (Church and Hanks, 1990), Poisson-Stirling (ps) (Church, 2000), Fisher's Exact Test (leftFisher, rightFisher, and twotailed) (Pedersen et al., 1996), Jaccard Coefficient (jaccard), Dice Coefficient (dice), Phi Coefficient (phi), t-score (tscore) (Church and Hanks, 1990), Pearson's Chi-Squared Test (x2), and the Odds Ratio (odds).

These measure the co-occurrence of word pairs (bigrams) relative to their individual frequencies and assess how likely it is that the word pair is occurring together by chance (and is therefore likely compositional) or has some significant pattern of occurrence as a pair (in which case it is non-compositional). More formally, many of these methods compare the observed empirical data with a model that casts the words in the bigram as independent statistical events. The measures determine the degree to which the observed data deviates from what would be expected under the model of independence. If the observed data differs significantly from that, then there is no evidence to support the hypothesis that the bigram is a chance event, and we assume that there is some interesting or significant pattern that implies non-compositionality. In some cases the training and test pairs are not adjacent (e.g., reinvent wheel for reinvent the wheel), and so window sizes of 2, 4, and 10 words were used when measuring the association between pairs of words. This means that 0, 2 and 8 intervening words were allowed, respectively.

Frequency count data for the word pairs are tabulated as shown in the example in Figure 1. The variable  $W_1$  represents the presence or absence of **red** in the first position of each word pair, and  $W_2$  represents the presence or absence of **tape** in the second position. This table tells us, for example, that *red tape* occurs 5,363 times  $(n_{11})$ , that *red* occurs 18,493 times  $(n_{1+})$ , and that bigrams that contain neither *red* nor *tape* occur 68,824,813 times  $(n_{22})$ . The total number of bigrams found in the corpus is  $68,845,263 (n_{++})$ . Note that these counts are based on a window size of 2. Counts increase with a larger window size. If the window size were 10, then  $n_{11}$  would tell us how many times *red* and *tape* occurred within 8 words of each other (in order).

$W_2$								
		tape	¬tape	totals				
$W_1$		$n_{11}=$	$n_{12} =$	$n_{1+}=$				
	red	5,363	13,130	18,493				
		n <sub>21</sub> =	n <sub>22</sub> =	$n_{2+}=$				
	$\neg red$	1,957	68,824,813	68,826,770				
		$n_{+1} =$	$n_{+2} =$	$n_{++}=$				
	totals	7,320	68,837,943	68,845,263				

Figure 1: Contingency Table Counts

#### 2.2 Scoring Word Pairs

The training pairs were ranked according to each of the measures in Text::NSP, where high scores indicate that two words ( $w_1$  and  $w_2$ ) are not occurring together by chance, and that there is a noncompositional meaning. However, high scores in the shared task meant exactly the opposite; that a word pair was highly compositional (and literal). In addition, the fine grained scoring in the shared task was on a scale of 0 to 100, and it was required that participating systems use that same scale. Thus, the scores from the measures were converted to this scale as follows:

Let the maximum value of the Text::NSP measure for all the pairs in the set under consideration be  $max(m(W_1, W_2))$ , where *m* represents the specific measure being used. Then the score for each word pair is normalized by dividing it by this maximum value, and subtracted from 1 and then multiplied by 100. More generally, the fine grained score for any word pair  $(w_1, w_2)$  as computed by a specific duluthx system is  $dx(w_1, w_2)$  and is calculated as follows:

$$dx(w_1, w_2) = 100 * \left(1 - \frac{m(w_1, w_2)}{max(m(W_1, W_2))}\right)$$
(1)

Coarse grained scoring is automatically performed by binning all of the resulting scores in the range 0-33 to *low*, 34 - 66 to *medium* and 67 - 100 to *high*.

Table 1: Text::NSP Rank Correlation with Gold St	andard
- duluth-1 corresponds to t-score window 10, du	uluth-2
with pmi window 10 and duluth-3 with pmi windo	ow 2

	Window Size				
Measure	2	4	10		
tscore	0.1484	0.2114	0.2674		
tmi	0.1335	0.1908	0.2361		
11	0.1336	0.1913	0.2358		
frequency	0.1865	0.2100	0.2126		
ps	0.0992	0.1554	0.1874		
x2	0.1157	0.1172	0.1654		
phi	0.1157	0.1167	0.1646		
jaccard	0.1253	0.1255	0.1602		
dice	0.1253	0.1255	0.1602		
odds	0.0216	0.0060	0.0257		
pmi	-0.0241	-0.0145	0.0143		
rightFisher	-0.1768	-0.0817	0.0740		
leftFisher	0.1316	0.0686	-0.0870		
twotailed	-0.1445	-0.0651	-0.1064		

### 2.3 Correlation of Word Pairs

Before the evaluation period, it was decided that duluth-1 (our flagship system) would be based on the measure of association that had the highest Spearman's rank correlation with the fine grained gold standard annotations of the training pairs. As can be seen from Table 1, that measure was the t-score based on a window size of 10.

As an additional experiment, the ranking of the training pairs according to each measure in Text::NSP was compared to the frequency ranking in the corpus. As can be seen in Table 2, once again it was the t-score that had the highest correlation.

While the correlation with the training pairs by the t-score was encouraging, the correlation with frequency was something of a surprise, and in fact caused some concern. Could a measure that correlated so highly with frequency really be successful in measuring semantic compositionality? However, upon reflection it seemed that correlation with frequency might be quite desirable, and led to the formulation of a second hypothesis:

Very frequent word pairs are more likely to be compositional (i.e., highly literal) than are less frequent word pairs.

Table 2: Text::NSP Rank Correlation with Frequency duluth-1 corresponds to t-score window 10, duluth-2 with pmi window 10 and duluth-3 with pmi window 2

	Window Size				
Measure	2	4	10		
tscore	0.9857	0.9578	0.8477		
ps	0.8856	0.8423	0.8299		
11	0.9082	0.8459	0.6953		
tmi	0.9080	0.8459	0.6951		
jaccard	0.7170	0.6128	0.5527		
dice	0.7170	0.6128	0.5527		
phi	0.7038	0.5743	0.4308		
x2	0.7039	0.5744	0.4303		
rightFisher	-0.5998	-0.3279	0.2004		
odds	0.3714	0.1483	-0.0353		
pmi	0.2487	0.0789	-0.1390		
leftFisher	0.5675	0.3500	-0.1726		
twotailed	-0.5965	-0.4434	-0.2712		

The assumption that underlies this hypothesis is that the most frequent word pairs tend to be very literal and non-compositional (e.g., *for the*, *in that*) and it would (in general) be a surprise to expect a compositional pair (e.g., *above board*, *rip saw*) to attain as high a frequency.

#### **3** duluth-1 (t-score in a 10 word window)

The duluth-1 system is based on the t-score in a 10 word window, and was selected because of its high correlation to the gold standard annotations of the training pairs and to the frequency ranking of the training pairs. The t-score optimizes both of our previous hypotheses, which suggests it should be a good choice for measuring compositionality.

By way of background, the t-score (t) is formulated as follows (Church et al., 1991), using the notation introduced in Figure 1 :

$$t = \frac{n_{11} - m_{11}}{\sqrt{n_{11}}} \tag{2}$$

where  $n_{11}$  is the observed count of the word pair, and  $m_{11}$  is the expected value based on the hypothesized model of independence between variables  $W_1$ and  $W_2$ . As such,

$$m_{11} = \frac{n_{1+} * n_{+1}}{n_{++}} \tag{3}$$

If there is little difference between the observed and expected values, then the t-score is closer to zero (or even less than zero) and the pair of words can be judged to occur together simply by chance (i.e., the hypothesis of independence is true).

The t-scores for the test pairs were converted following equation (1), and then submitted for evaluation. duluth-1 placed in the middle ranks in the fine grain evaluation according to mean distance, and was the top ranked system according to the label precision evaluation of coarse grained scoring.

### 4 duluth-2 (pmi with window size of 10)

In studying Tables 1 and 2, it's clear that Pointwise Mutual Information (pmi) deviates rather significantly from frequency and the t-score. At the time of the evaluation, we did not know if our hypotheses that motivated the use of the t-score would prove to be true. If they did not, it seemed sensible to include the most opposite measure to the t-score, as a kind of fail safe mechanism for our systems overall. In addition, pmi has a fairly significant history of use in identifying collocations and features for other NLP tasks (e.g., (Pantel and Lin, 2002)), and so it seemed like a credible candidate.

pmi has a well known bias towards identifying words that only occur together, and tends to prefer less frequent word pairs, and this is why it diverges so significantly from the t-score and frequency. Interestingly, pmi is also based on the same observed and expected values  $n_{11}$  and  $m_{11}$  as used in the tscore (and many of the other measures), and is calculated as follows:

$$pmi = log \frac{n_{11}}{m_{11}}$$
 (4)

If there is little difference between the observed and expected values, then pmi tends towards 0 and we treat the word pairs as independent and compositional.

duluth-2 relies on a window size of 10, since it diverges dramatically from the t-score and frequency.

## 5 duluth-3 (pmi with window size of 2)

duluth-3 is a very close relative of duluth-2, and differs only in that it requires word pairs to be adjacent. Given the wider window sizes in duluth-2, it is clear that if a pair has a high pmi score, they must only occur (mostly) together. duluth-3 only considers adjacent words, and so the words that make up the pairs may also appear elsewhere in the corpus. As such duluth-3 may tend to assign higher pmi scores than the more exacting duluth-2 (where high scores mean low compositionality). And in fact this is what occurred. In the coarse scoring scheme, duluth-1 only identified 2 low compositional word pairs, whereas duluth-2 identified 46 and duluth-3 identified 70.

Despite the difference in the window size the rank correlation between duluth-2 and duluth-3 is relatively high (.9330). Both performed comparably in the evaluation, being near the bottom of both the fine and coarse grained evaluations. By comparison, duluth-1 and duluth-2 have a relatively low rank correlation of .1756, and duluth-1 and duluth-3 have a modest correlation of .3438.

## 6 Conclusions

The Duluth systems seek to evaluate the degree to which measures of collocation are able to measure semantic compositionality as well. The results of this shared task suggest that the t-score is well suited to make coarse grained distinctions between high, medium, and low levels of compositionality, since duluth-1 was the top ranked system in the coarse grained evaluation. While this success might be considered surprising due to the simplicity of the approach, it should not be underestimated. There are two separate hypotheses that underly the t-score and its use in measuring semantic compositionality. These hold that word pairs with high measures of association are more likely to be non-compositional, and that more frequent word pairs are more likely to be compositional. Of the measures evaluated in this study, the t-score was best able to optimize both of these conditions.

## 7 Acknowledgements

The experiments in this paper were conducted with version 1.23 of the Ngram Statistics Package (Text::NSP), which is implemented in Perl and freely available from http://ngram.sourceforge.net.

## References

- S. Banerjee and T. Pedersen. 2003. The design, implementation, and use of the Ngram Statistics Package. In Proceedings of the Fourth International Conference on Intelligent Text Processing and Computational Linguistics, pages 370–381, Mexico City, February.
- M. Baroni, S. Bernardini, A. Ferraresi, and E. Zanchetta. 2009. The WaCky wide web: A collection of very large linguistically processed web-crawled corpora. *Language Resources and Evaluation*, 43(3):209–226.
- C. Biemann and E. Giesbrecht. 2011. Distributional semantics and compositionality 2011: Shared task description and results. In *Proceedings of DiSCo–2011 in conjunction ACL HLT 2011*, Portland, Oregon, June. Association for Computational Linguistics.
- K. Church and P. Hanks. 1990. Word association norms, mutual information and lexicography. In *Proceed*ings of the 28th Annual Meeting of the Association for Computational Linguistics, pages 76–83.
- K. Church, W. Gale, P. Hanks, and D. Hindle. 1991. Using statistics in lexical analysis. In U. Zernik, editor, *Lexical Acquisition: Exploiting On-Line Resources to Build a Lexicon*. Lawrence Erlbaum Associates, Hillsdale, NJ.
- K. Church. 2000. Empirical estimates of adaptation: The chance of two noriegas is closer to p/2 than p2. In Proceedings of the 18th International Conference on Computational Linguistics (COLING-2000), pages 180–186, Saarbrücken, Germany.
- T. Dunning. 1993. Accurate methods for the statistics of surprise and coincidence. *Computational Linguistics*, 19(1):61–74.
- P. Pantel and D. Lin. 2002. Discovering word senses from text. In Proceedings of ACM SIGKDD Conference on Knowledge Discovery and Data Mining-2002.
- T. Pedersen, M. Kayaalp, and R. Bruce. 1996. Significant lexical relationships. In *Proceedings of the Thirteenth National Conference on Artificial Intelligence*, pages 455–460, Portland, OR, August.